

Methods for Noise Cancellation based on the EM Algorithm ¹

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Simplified scenario

• Demonstration 1:

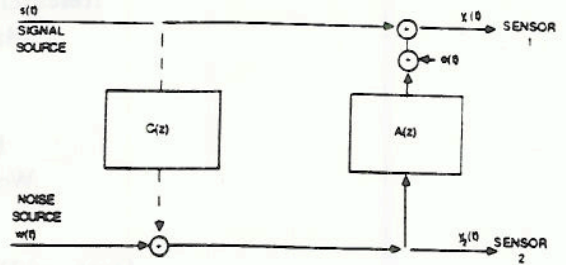
$C(z) = 0$. $A(z)$: 10^{th} order FIR filter.
SNR for sensor 1: -20 to -25 dB.

- Noisy signal (sensor 1)
- Enhanced signal (EM algorithm, version 1)
- Original (clean) signal

• Demonstration 2:

$C(z) = 0.1z^{-5}$. $A(z)$: 10^{th} order FIR filter.
SNR for sensor 1: -20 to -25 dB.

- Noisy signal (sensor 1)
- Enhanced signal (least-squares method)
- Enhanced signal (EM algorithm, version 1)
- Original (clean) signal



More general scenario

• Demonstration 3:

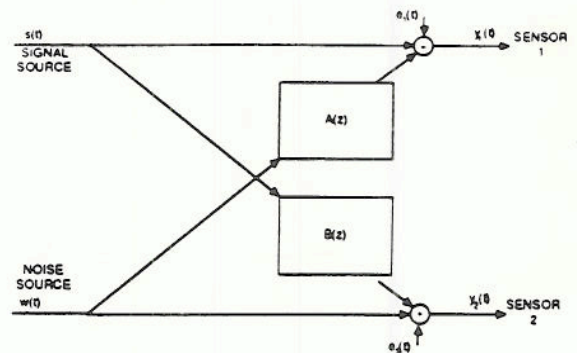
$A(z), B(z)$: 10^{th} order FIR filters.
SNR for sensor 1: -20 to -25 dB.

- Noisy signal (sensor 1)
- Enhanced signal (least-squares method)
- Enhanced signal (EM algorithm, version 2)

• Demonstration 4:

$A(z), B(z)$: realistic room impulse responses.
SNR for sensor 1: +20 dB.

- Measured signal (sensor 1)
- Enhanced signal (least-squares method)
filter order 250
- Enhanced signal (EM algorithm, version 2)
filter order 250



• Demonstration 5:

$A(z), B(z)$: realistic room impulse responses.
SNR for sensor 1: 0 dB.

- Measured signal (sensor 1)
- Enhanced signal (least-squares method)
filter order 500
- Enhanced signal (EM algorithm, version 2)
filter order 250
- Enhanced signal (EM algorithm, version 2)
filter order 500

PROBLEM FORMULATION

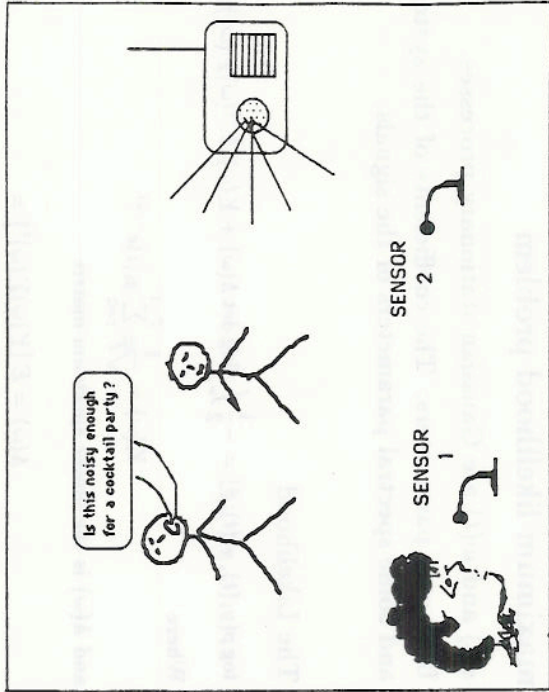
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MAIN IDEA

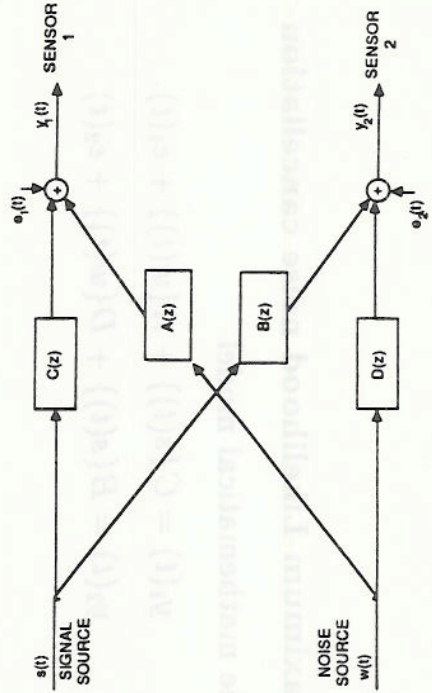
MAXIMUM LIKELIHOOD APPROACH

vs.

LEAST SQUARES APPROACH



The Model:



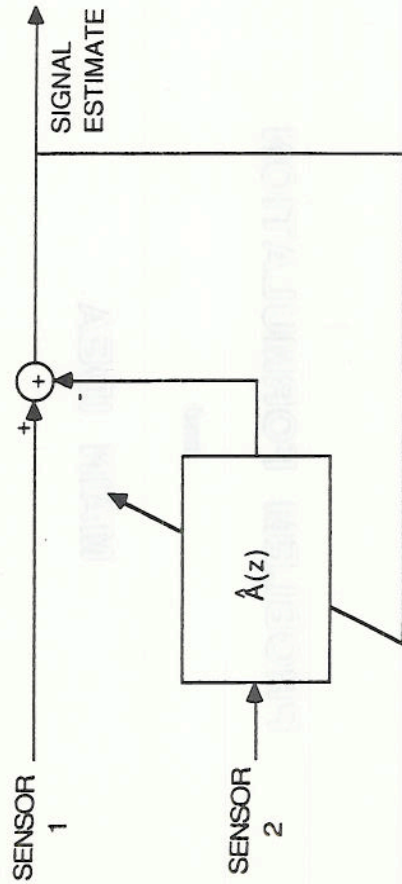
The least-squares solution (Widrow)

- Estimate the coefficients $\{a_k\}$ of a system $A(z)$, by least-square fitting of $y_2(t)$ (the reference noise signal) to $y_1(t)$ (the primary speech plus noise signal).

- Cancel the noise;

$$\hat{s}(t) = y_1(t) - \sum_{k=0}^p \hat{a}_k y_2(t - k)$$

- The least-square solution is approximated by a recursive algorithm; LMS (Widrow), or RLS.



Maximum Likelihood noise cancellation

- The mathematical model

$$y_1(t) = C\{s(t)\} + A\{w(t)\} + e_1(t)$$

$$y_2(t) = B\{s(t)\} + D\{w(t)\} + e_2(t)$$

- A maximum likelihood problem

- $s(t)$ and $w(t)$ are Gaussian stationary processes.
- Unknown parameters: The coefficients of the systems and some spectral parameters of the signals.

- The Likelihood

$$\log p(y_1(t), y_2(t); \theta) = -\frac{1}{2} \int_{\omega} (\log \det \Lambda(\omega) + Y(\omega)^T \Lambda^{-1}(\omega) Y(\omega)) d\omega$$

Where

$$Y_i(\omega) = \frac{1}{\sqrt{T}} \sum_{t=0}^{T-1} y_i(t) e^{-j\omega t}$$

and $\Lambda(\omega)$ is the power spectrum matrix

$$\Lambda(\omega) = E \{ Y(\omega) Y(\omega)^H \} =$$

$$\begin{bmatrix} C(\omega)P_s(\omega)C^*(\omega) + A(\omega)P_w(\omega)A^*(\omega) + \sigma_s^2 & C(\omega)P_s(\omega)B^*(\omega) + A(\omega)P_w(\omega)D^*(\omega) \\ B(\omega)P_s(\omega)C^*(\omega) + D(\omega)P_w(\omega)A^*(\omega) & B(\omega)P_s(\omega)B^*(\omega) + D(\omega)P_w(\omega)D^*(\omega) + \sigma_s^2 \end{bmatrix}$$

Solution, using the EM approach

- Given $A(z)$ and the spectral parameters of $s(t) \implies$ Estimate $s(t)$ and $n(t)$.
- Having $s(t), n(t)$ (and $y_2(t)$) \implies
 - Get the maximum likelihood estimate of $\{a_k\}$ and σ^2 by least-square fitting of $y_2(t)$ to $n(t)$.
 - Get the ML estimate of the spectral parameters of $s(t)$.

The Algorithm

- **The E step**
having a current estimate of the parameters
 - Generate a signal $x(t)$

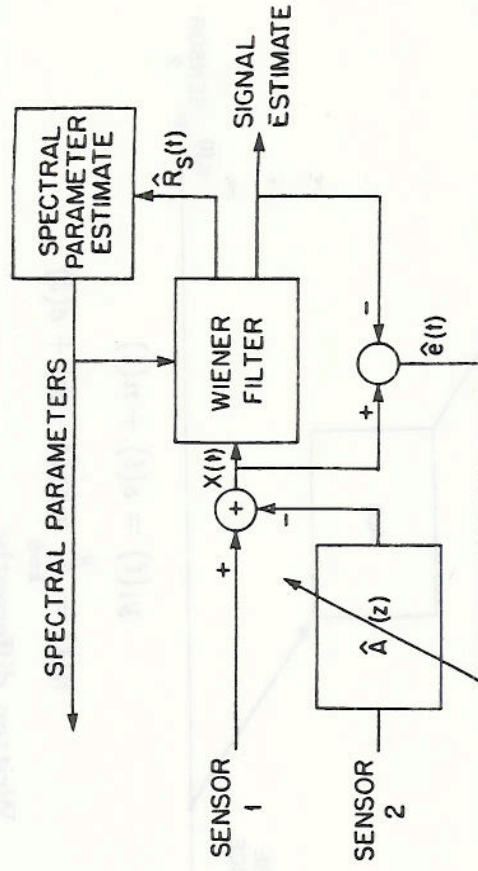
$$x(t) = y_1(t) - \sum_{k=0}^p a_k y_2(t-k)$$
- (Note that given $\{a_k\}$, $x(t) = s(t) + \epsilon(t)$.)
- Apply a Wiener filter to $x(t)$.
 - Get an estimate of $\epsilon(t), s(t)$ and of the sample covariance of $s(t)$.

- **The M step**
 - Solve the following least squares problem

$$\min_{\{a_k\}} \sum_t (\epsilon(t) - \sum_{k=0}^p a_k y_2(t-k))^2$$
 - Add the result to the previous estimate of $\{a_k\}$ to get a new estimate of $\{a_k\}$.
 - Update the signal spectral parameter
For LPC parameters, solve the normal equation using the estimated sample covariance matrix.

The suggested processing scheme

(EM algorithm, version 1)



MAXIMUM LIKELIHOOD NOISE CANCELLATION

via

THE E-M ALGORITHM

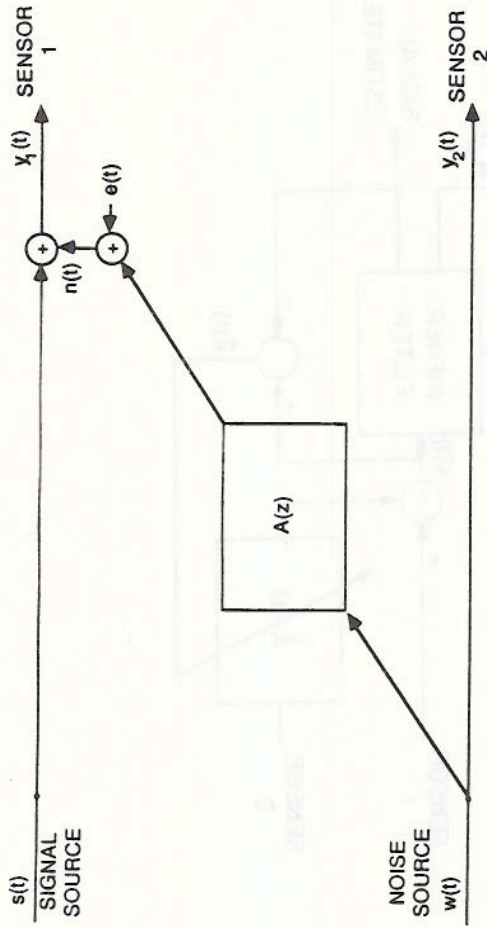
A

SIMPLIFIED

SCENARIO

Noise cancellation, simplified scenario.

Observe $y_1(t)$ and $y_2(t)$,



$$y_1(t) = s(t) + n(t)$$

$$n(t) = \sum_{k=0}^p a_k y_2(t - k) + e(t)$$

Written differently,

$$y_1(t) = s(t) + \sum_{k=0}^p a_k w(t - k) + e(t)$$

$$y_2(t) = w(t)$$

MAXIMUM LIKELIHOOD NOISE CANCELLATION

via

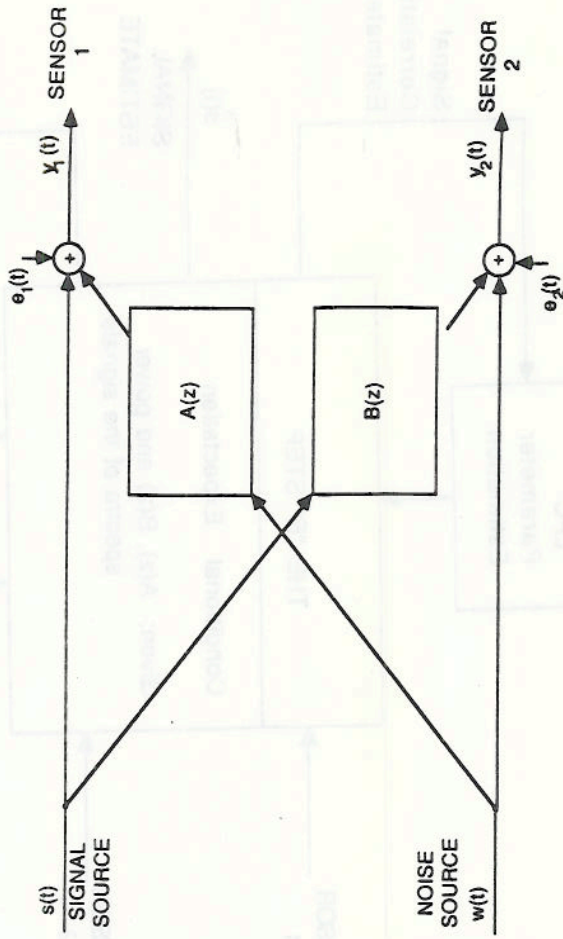
THE EM ALGORITHM

A

MORE GENERAL

SCENARIO

Two microphones, more general situation



$$y_1(t) = s(t) + A\{w(t)\} + e_1(t) = s(t) + n_1(t)$$

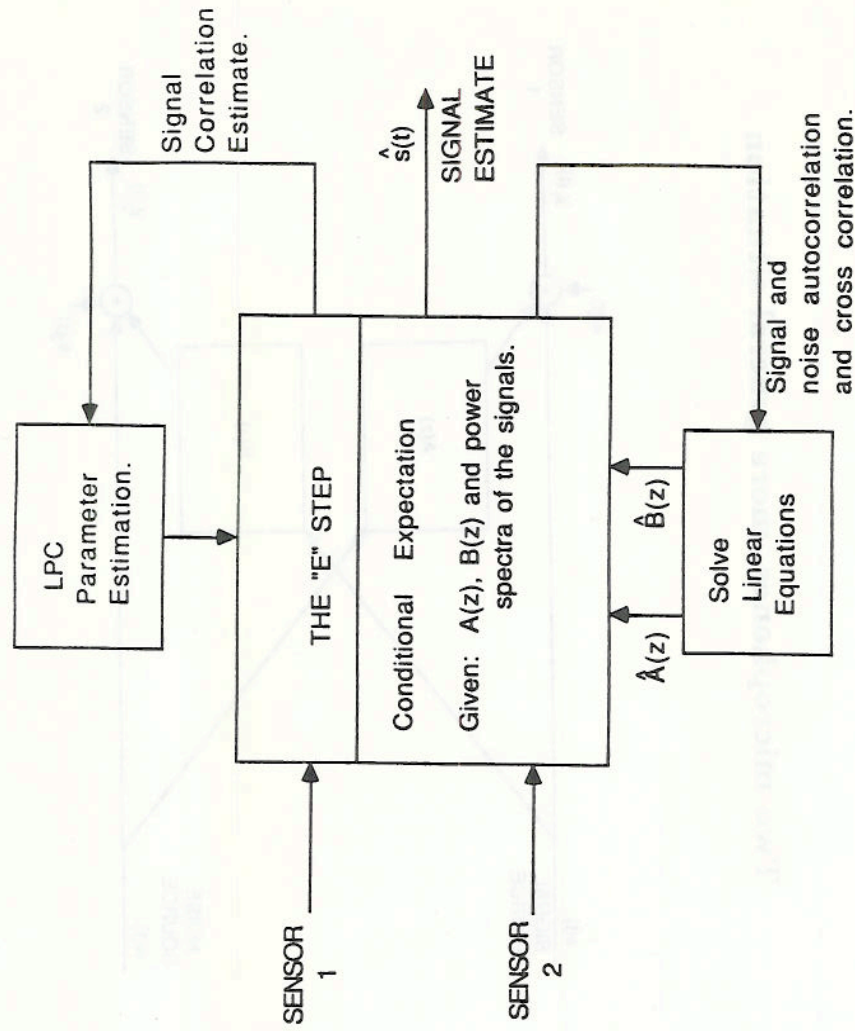
$$y_2(t) = B\{s(t)\} + w(t) + e_2(t) = w(t) + n_2(t)$$

Solution, using the EM approach

- The complete data is the outputs *and* inputs of the systems $A(z), B(z)$.
- Given $A(z), B(z)$ and the spectral parameters of $s(t), w(t) \implies$ Estimate the complete data and its various quadratic terms. **(E step)**
- Having the complete data and its quadratic terms \implies
 - Get the maximum likelihood estimate of $A(z), B(z)$ coefficients by solving linear equations.
 - Get the ML estimate of the spectral parameters. **(M step)**

The suggested processing scheme

(EM algorithm, version 2)



The EM algorithm: Basic theory

- Maximizing the likelihood of the observations $\log p(\underline{y}; \theta)$ is complicated
- The observations (\underline{y}) , can be viewed as “incomplete”. The “complete data” \underline{x} contains the observations as:

$$\underline{y} = g(\underline{x}), \quad g \text{ non-invertible}$$

- If we had observed \underline{x} we would estimate θ easily, i.e

$$\max_{\theta} \log p_{\underline{x}}(\underline{x}; \theta) \quad \text{is easy!}$$

The Algorithm

- Start at $\theta^{(0)}$
- Iterate:

1. E step:

$$Q(\theta, \underline{\theta}^{(k)}) = E \left\{ \log p(\underline{x}; \theta) / \underline{y}; \underline{\theta}^{(k)} \right\}$$

For exponential family, estimate the sufficient statistics $E \{ t(\underline{x}) / \underline{y}; \underline{\theta}^{(k)} \}$ and substitute in the expression for the Likelihood of \underline{x} .

2. M step:

$$\underline{\theta}^{(k+1)} \leftarrow \max_{\theta} Q(\theta, \underline{\theta}^{(k)})$$

For exponential family, $Q(\theta, \underline{\theta}^{(k)})$ has the same functional form as $\log p(\underline{x}; \theta)$

The Estimate Maximize (EM) Algorithm¹

The Likelihood of the observations is related to the Likelihood of the "complete Data":

$$p(\underline{y}; \theta) = \frac{p(\underline{x}, \underline{y}; \theta)}{p(\underline{x}/\underline{y}; \theta)}$$

But $\underline{y} = g(\underline{x})$ is given. Define the set $\mathcal{X}(\underline{y})$

$$\mathcal{X}(\underline{y}) = \{\underline{x} \mid \underline{y} = g(\underline{x})\}$$

For any $\underline{x} \in \mathcal{X}(\underline{y})$ $p(\underline{x}, \underline{y}; \theta) = p(\underline{x}; \theta)$

or:

$$p(\underline{y}; \theta) = \frac{p(\underline{x}; \theta)}{p(\underline{x}/\underline{y}; \theta)}$$

$$\log p(\underline{y}; \theta) = \log p(\underline{x}; \theta) - \log p(\underline{x}/\underline{y}; \theta) \quad (1)$$

We can substitute in the R.H.S of (1) any $\underline{x} \in \mathcal{X}(\underline{y})$ or we can average it over this set and still have an expression for the likelihood of the observations.

Suppose we average the R.H.S of (1) using

$$p(\underline{x}/\underline{y}; \underline{\theta}') \quad \underline{\theta}' \text{ arbitrary}$$

$$\underbrace{\log p(\underline{y}; \underline{\theta})}_{L(\underline{\theta})} = \underbrace{E \{ \log p(\underline{x}; \underline{\theta}) / \underline{y}; \underline{\theta}' \}}_{Q(\underline{\theta}, \underline{\theta}')} - \underbrace{E \{ \log p(\underline{x}/\underline{y}; \underline{\theta}) / \underline{y}; \underline{\theta} \}}_{H(\underline{\theta}, \underline{\theta}')}$$

- $H(\underline{\theta}, \underline{\theta}')$ is in the form $E_f \log g$
- $H(\underline{\theta}', \underline{\theta}')$ is in the form $E_f \log f$ f, g PDF

Since (by Jensen's Inequality)

$$E_f \log f \geq E_f \log g \quad \text{i.e. } H(\underline{\theta}', \underline{\theta}') \geq H(\underline{\theta}, \underline{\theta}')$$

Then if we find $\underline{\theta}$ s.t

$$Q(\underline{\theta}, \underline{\theta}') \geq Q(\underline{\theta}', \underline{\theta}')$$

we have

$$L(\underline{\theta}) \geq L(\underline{\theta}')$$

¹ A.P Dempster, N.M Laird and D.B Rubin "Maximum Likelihood from incomplete data via the EM Algorithm", *J. Royal Stat. Soc.* no. 1 pp.1-38, Jan 1977